A Model of the Fed's View on Inflation

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"Inflation is characterized by an underlying trend that has been essentially constant since the mid-1990s; Theory and evidence suggest that this trend is strongly influenced by inflation expectations that, in turn, depend on monetary policy. In particular, the remarkable stability of various measures of expected inflation in recent years presumably represents the fruits of the Federal Reserve's sustained effort since the early 1980s to bring down and stabilize inflation at a low level. The anchoring of inflation expectations ... does not, however, prevent actual inflation from fluctuating from year to year in response to the temporary influence of movements in energy prices and other disturbances. In addition, inflation will tend to run above or below its underlying trend to the extent that resource utilization-which may serve as an indicator of firms' marginal costs-is persistently high or low."

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Challenges to this view:

- Weak empirical evidence on the PC
- Inflation can be forecast by statistical processes unrelated to slack
- Evidence of the **flattening** (or **disappearance** of the Phillips Curve)
- Missing deflation...
- Disanchoring of consumers' expectation due to oil shocks

This Paper

An Econometric Formalisation of the Policymakers'/Median Economist's View:

- A semi-structural unobserved components Trend-Cycle model á la Harvey (1985)
- ... employing survey data on inflation
- ... encompasses full information rational expectations (FIRE) but allows for deviations
- **Bayesian estimation** (Harvey et al., 2007, Del Negro et al. 2017, and Lenza and Jarociński, 2018)

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Results:

- 1. Stable expectational trend
- 2. Sizeable and fairly steep reduced form Phillips curve
- 3. Some rationalisation of the **inflation puzzles**

An Empirical Trend-Cycle Model of Inflation Dynamics

A Stylised Rational Expectations Model

$$\begin{pmatrix} y_t \\ \pi_t \\ \mathbb{E}_t \left[\pi_{t+1} \right] \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ \delta_{\pi} & 1 \\ \delta_{exp,1} + \delta_{exp,2}L & 1 \end{pmatrix} \begin{pmatrix} \widehat{\psi}_t \\ \mu_t^{\pi} \end{pmatrix} + \begin{pmatrix} \mu_t^y \\ 0 \\ 0 \end{pmatrix} + \begin{pmatrix} \psi_t^y \\ \psi_t^{\pi} \\ 0 \end{pmatrix}$$

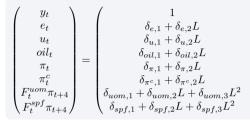
• Can accommodate different specifications for the Phillips Curve

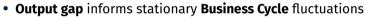
$$\hat{\pi}_t = \sum_{i=1}^2 \delta_i \hat{\pi}_{t-i} + \beta \mathbb{E}_t \left[\hat{\pi}_{t+1} \right] + \kappa \hat{y}_t + v_t$$

- AR(2) $\widehat{\psi}_t$ with complex roots would be a solution to hybrid Phillips Curve
- AR(1) $\widehat{\psi}_t$ would be the solution to a **purely forward** looking NK Phillips Curve
- It also nests the backwards looking 'Old-Keynesian' Phillips curve connecting output gap and prices

A Richer Inflation Dynamics

- ① Heterogenous dynamics along the business cycle
 ⇒ Lagged relations prices-slack
- ② Labour market dynamics along the business cycle
 ⇒ Okun's law connecting slack-unemployment
- 3 Energy price can impact CPI
 - as markup shocks
 - directly as consumption good
 - via expectations (non-fundamental fluctuations)
- Deviations from full-information RE





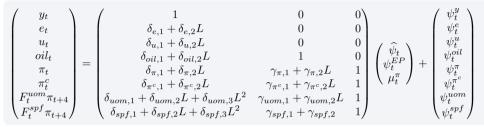
- ... connects to labour market variables via Okun's law
- ... and to prices and expectations via the Phillips curve

$$\begin{pmatrix} y_t \\ e_t \\ u_t \\ oil_t \\ \pi_t \\ \pi_t^c \\ F_t^{uom} \pi_{t+4} \\ F_t^{spf} \pi_{t+4} \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ \delta_{e,1} + \delta_{e,2}L & 0 \\ \delta_{oil,1} + \delta_{oil,2}L & 0 \\ \delta_{oil,1} + \delta_{oil,2}L & 0 \\ \delta_{\pi,c} + \delta_{\pi^c,2}L & 1 \\ \delta_{\pi^c,1} + \delta_{\pi^c,2}L & 1 \\ \delta_{uom,1} + \delta_{uom,2}L + \delta_{uom,3}L^2 & 1 \\ \delta_{spf,1} + \delta_{spf,2}L + \delta_{spf,3}L^2 & 1 \end{pmatrix} \begin{pmatrix} \psi_{m} \\ \psi_{m} \\ \psi_{m} \end{pmatrix}$$

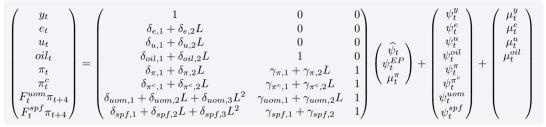
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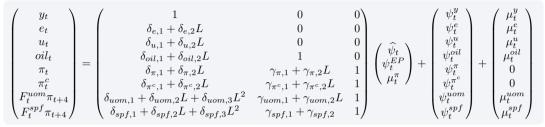
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- Stationary idiosyncratic disturbances

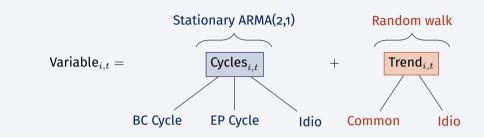


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- Independent trend in employment/unemployment (trend employment/equilibrium unemployment)



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- ... and idiosyncratic trends in expectations

The Model in a Nutshell



• Stationary Cycles

$$\begin{pmatrix} \psi_t^j \\ \psi_t^{*j} \end{pmatrix} = \rho^j \begin{pmatrix} \cos(\lambda^j) & \sin(\lambda^j) \\ -\sin(\lambda^j) & \cos(\lambda^j) \end{pmatrix} \begin{pmatrix} \psi_{t-1}^j \\ \psi_{t-1}^{*j} \end{pmatrix} + \begin{pmatrix} v_t^j \\ v_t^{*j} \end{pmatrix}, \quad \begin{pmatrix} v_t^j \\ v_t^{*j} \end{pmatrix} \sim \mathcal{N}(0,\varsigma_j^2 I_2)$$

• Unit Root Trends (w/ or w/o drift)

$$\mu_t^j = \mu_0^j + \mu_{t-1}^j + u_t^j, \quad u_t^j \sim \mathcal{N}(0,\sigma_j^2) \;.$$

Bringing the Model to the Data

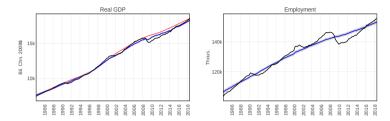
Variable	Transformation	Loads on		
		BC Cycle	EP Cycle	Common Trend
Gross Domestic Product	Levels	\checkmark	×	×
Employment (or Empl/Pop)	Levels	\checkmark	×	×
Unemployment Rate	Levels	\checkmark	×	×
WTI Spot Oil Price	Levels	\checkmark	\checkmark	×
CPI: All Items	YoY	\checkmark	\checkmark	\checkmark
Core CPI	YoY	\checkmark	\checkmark	\checkmark
UoM: Expected Inflation	Levels	\checkmark	\checkmark	\checkmark
SPF: Expected Inflation	Levels	\checkmark	\checkmark	\checkmark

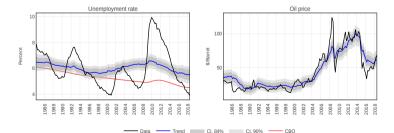
Sample: Quarterly, Q1-1984 to Q2-2017

Bayesian Estimation

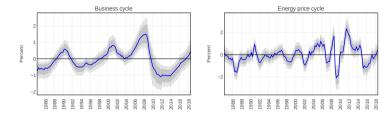
Trends & Cycles in US Inflation

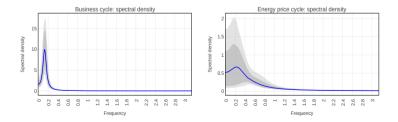
Output Potential, Equilibrium Employment





Common Cycles

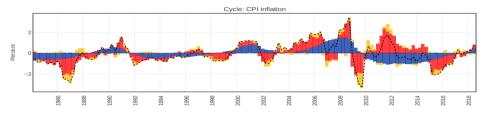




Median CI, 90% CI, 84%

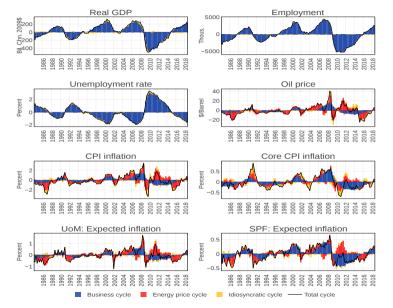
Historical Decomposition of the CPI

Headline CPI





Common Cycles



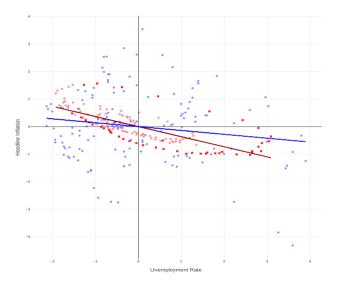
Output Gap

How Deep was Last Recession?

Percent Greenbook — CBO — Trend-Cycle Model

Output gap as a percentage of potential GDP

The Slope of the Phillips Curve

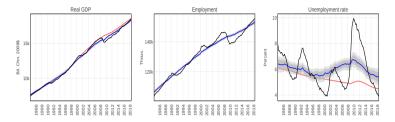


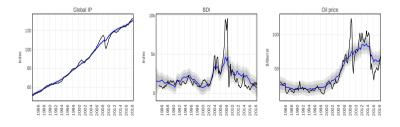
Phillips Curve slope

- Blue line is -0.14
- Red line is -0.39

Global Determinants of Inflation

Global: Output Potential, Equilibrium Employment

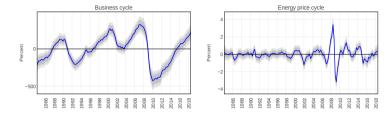


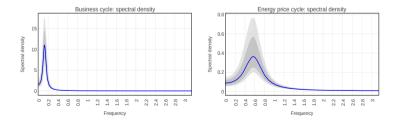


Trend Cl. 84% Cl. 90% CBO

15/22

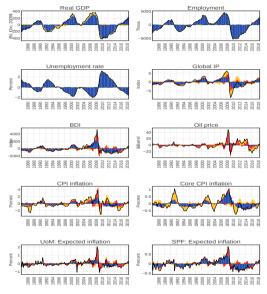
Global: Output Potential, Equilibrium Employment





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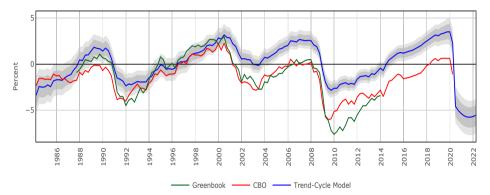
Global: Historical Decomposition



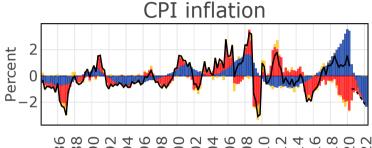
A Rough & Ready COVID Exercise

COVID Output Gap

Output gap as a percentage of potential GDP

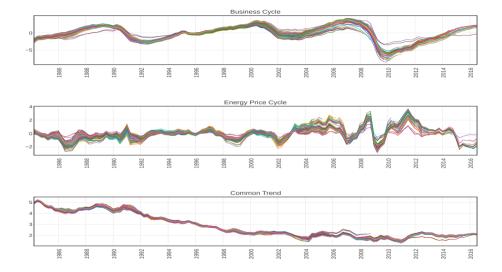


Deflation Ahead?



Out-of-Sample Performances

Out-of-Sample Cycle & Trend Revisions



Out-of-Sample Forecast Evaluation

Root Mean Squared Forecast Error relative to the Random Walk with drift

Horizon	Variable	TC Model	BVAR	UC-SV	Horizon	Variable	TC Model	BVAR	UC-SV
h=1	Real GDP	1.04	0.93	x	h=4	Real GDP	1.11	1.01	х
	Employment	0.98	0.76	х		Employment	1.06	0.82	х
	Unemployment rate	0.85	0.67	х		Unemployment rate	0.86	0.83	х
	Oil price	1.03	1.08	х		Oil price	1.03	1.26	х
	CPI Inflation	0.94	0.91	1.00		CPI Inflation	0.87	1.13	0.97
	Core CPI Inflation	1.01	1.04	1.01		Core CPI Inflation	0.95	1.22	0.96
	UOM: Expected inflation	0.98	1.04	х		UOM: Expected inflation	0.96	1.14	х
	SPF: Expected CPI	0.95	1.06	х		SPF: Expected CPI	0.92	1.31	х
h=2	Real GDP	1.06	0.93	x		Real GDP	1.17	1.21	х
	Employment	1.00	0.76	х	h=8	Employment	1.13	1.01	х
	Unemployment rate	0.85	0.71	х		Unemployment rate	0.85	1.02	х
	Oil price	1.04	1.18	х		Oil price	0.99	1.36	х
	CPI Inflation	0.90	0.98	0.99		CPI Inflation	0.81	1.09	0.95
	Core CPI Inflation	0.99	1.15	0.99		Core CPI Inflation	0.84	1.30	0.91
	UOM: Expected inflation	0.98	1.09	х		UOM: Expected inflation	0.92	1.28	х
	SPF: Expected CPI	0.94	1.18	х		SPF: Expected CPI	0.88	1.34	х

Conclusions

- Explicit modelling of **trends & cycles** and **parsimonious** characterisation of the **structural** relationships amongst macro variables
- Inflation dynamics since the 1980s shows a stable and fairly steep reduced form **Phillips Curve** with maximum power at around eight years periodicity.
- The Phillips Curve is **not** always the **dominant component** of cyclical inflation
- Long-term inflation expectations common trend between inflation and expectations roughly stable
- Consumer survey shows large and persistent deviations from the common trend (Coibion and Gorodnichenko, 2015)
- Important question on what happened to **trend growth/output gap** (Fernald et al. 2017, Coibion et al. 2018, Blanchard et al., 2015)

Appendix

Priors

Table: Prior distributions

Name	Support	Density	Parameter 1	Parameter 2
δ , γ , ϕ and $ au$	${\rm I\!R}$	Normal	0	1000
σ^2 and $arsigma^2$	$(0,\infty)$	Inverse-Gamma	3	1
ho	[0.001, 0.970]	Uniform	0.001	0.970
λ	$[0.001,\pi]$	Uniform	0.001	π



Metropolis-Within-Gibbs Algorithm

The algorithm is structured in two blocks

- The **first block** uses a Metropolis step for the **estimation of the state-space parameters**
- The **second block** uses a Gibbs step to draw the **unobserved states** conditional on the model parameters



- Metropolis algorithm draws the model parameters in the unbounded space in order to avoid a-priori rejections and to obtain a more efficient estimation routine
- The following transformations have been applied to parameters with Normal, Inverse-Gamma and Uniform priors

$$\theta_j^N = \Theta_j^N \qquad \theta_j^{IG} = \ln(\Theta_j^{IG} - a_j) \qquad \theta_j^U = \ln\left(\frac{\Theta_j^U - a_j}{b_j - \Theta_j^U}\right)$$

Where a_j and b_j are the lower and the upper bounds for the *j*-th parameter

• Jacobians of the transformations of the variables

$$\ln\left(\frac{d\Theta_j^N}{d\theta_j^N}\right) = 0 \qquad \ln\left(\frac{d\Theta_j^{IG}}{d\theta_j^{IG}}\right) = \theta_j^{IG}$$
$$\ln\left(\frac{d\Theta_j^U}{d\theta_j^U}\right) = \ln(b_j - a_j) + \theta_j^U - 2\ln(1 + \exp(\theta_j^U))$$



Algorithm: Metropolis-Within-Gibbs Initialisation

For $s = 1, ..., n_s$ ($n_s = 40000$)

- 1. Metropolis Algorithm
 - i. Draw a candidate vector for the unbounded parameters (θ_*), from a multivariate normal distribution with mean θ_{s-1} and variance $\omega \mathbb{I}$, where ω is a scaling constant used to get an acceptance rate between 25% and 35%

ii. Set

$$heta_s = egin{cases} heta_* & ext{with probability } \eta \ heta_{s-1} & ext{with probability } 1-\eta \end{cases}$$
 (1)

for

$$\eta = \min\left(1, \frac{p(y \mid f(\theta_*)^{-1}) \, p(f(\theta_*)^{-1}) \, J(\theta_*)}{p(y \mid f(\theta_{s-1})^{-1}) \, p(f(\theta_{s-1})^{-1}) \, J(\theta_{s-1})}\right) \tag{2}$$

2. Discard the first $s = 1, \ldots, n_0$ ($n_0 = 20000$) draws of θ_s .

Back...

Algorithm: Metropolis-Within-Gibbs

Recursion

1. Metropolis Algorithm

Set Σ to the sample covariance of the chain of θ_s , ($s = \{n_0, \ldots, n_s\}$), from the Initialisation step.

For $q = 1, \dots, n_q$ ($n_q = 20000$)

i. Draw a candidate vector for the parameters (θ_*), from a multivariate normal distribution with mean θ_{q-1} and variance $\omega\Sigma$, where ω is set to have an acceptance rate between 25% and 35%

ii. Set

$$\theta_q = \begin{cases} \theta_* & \text{with probability } \eta \\ \theta_{q-1} & \text{with probability } 1 - \eta \end{cases}$$
(1)

where η is defined as in the Initialisation step.

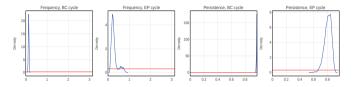
2. Gibbs sampling

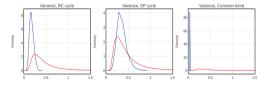
For $n_q > n_{\emptyset}$ for $n_{\emptyset} = 10000$ (burn-in period), apply the univariate approach for multivariate time series of of Koopman and Durbin (2000) to the simulation smoother proposed in Durbin and Koopman (2002) to sample the unobserved states, conditional on the parameters. In doing so, we follow the refinement proposed in Jarociński (2015).

3. Discard the first $q = 1, \ldots, n_{\emptyset}$ draws of θ_q .



Priors and Posteriors Variance of Shocks to the Components





Prior Posterior

