The Impact of Debt Relief Generosity and Liquid Wealth on Household Bankruptcy

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Consumer Bankruptcy in the US

- Bankruptcy is a major source of debt relief
 - 1 million households file each year
 - ▶ 1/10 Americans have filed at some point in their life
 - ► Transfers 3× the resources to households as unemp. insurance
- Trade-offs of generous bankruptcy
 - ► Can create moral hazard ⇒ discourages lending
 - ► Helps smooth consumption \Rightarrow provides insurance

• Focus: importance of moral hazard vs. incomplete insurance

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- Analysis I: impact of debt relief generosity
 - Approach: regression kink design (RKD) using kink in generosity due to exemption laws

- Analysis II: impact of mortgage payment reductions
 - ► Approach: IV strategy using variation in contract feature of adjustable-rate mortgages

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 - Fixes wealth out of bankruptcy, varies the wealth gain from filing
 - Isolates a "strategic" default motive
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 - ► Approach: IV strategy using variation in contract feature of adjustable-rate mortgages
 - Fixes wealth gain from filing, varies wealth in and out of bankruptcy
 - ▶ Isolates a "cash-flow" default motive

Main Findings

Empirical Results

- \$1,000 reduction in generosity leads to a 3% fall in filings
- ▶ \$1,000 reduction in payments leads to a 12% fall in filings

Model Implications

- ▶ Relatively stronger cash-flow motive ⇒ strong desire to avoid bankruptcy
- "Other" costs of bankruptcy are large (i.e., stigma, credit market exclusion)

Key Conclusions

- Moral hazard is a weak driver of bankruptcy
- Incomplete insurance is a strong driver of bankruptcy
- Suggests welfare-improving scope for generous bankruptcy

Related Literature & Contributions

Strategic Default: Elul et al. (2010); Guiso, Sapienza, and Zingales (2013); Mayer, Morrison, Piskorski, and Gupta (2014); Eberly and Krishnamurthy (2014); Gerardi, Herkenhoff, Ohanian, and Willen (2017); Bhutta, Dokko, and Shan (2017); Ganong and Noel (2018); Dobbie and Song (2018); Gupta and Hansman (2018); Hsu, Matsa, and Melzer (2018)

Trade-Offs of Generous Bankruptcy: Li and Sarte (2006); Livshits, MacGee, and Tertilt (2007); Chatterjee, Corbae, Ríos-Rull, and Nakajima (2007); Elul and Gottardi (2015); Mitman (2016); Dávila (2016); Auclert, Dobbie, and Goldsmith-Pinkham (2019); Gross, Kluender, Liu, Notowidigdo, and Wang (2019); Auclert and Mitman (2019)

Causes of Bankruptcy Fay, Hurst, and White (2002); Gross and Notowidigdo (2011); Gross, Notowidigdo, and Wang (2014); Mahoney (2015); Keys (2018); Kleiner, Stoffman, and Yonker (2019)

Effects of Mortgage Payment Reductions: Di Maggio, Kermani, Keys, Piskorski, Ramcharan, Seru, and Yao (2017); Fuster and Willen (2017); Campbell, Clara, and Cocco (2018); Gupta (2019)

Regression Kink Designs: Calonico, Cattaneo, and Titiunik (2014); Card, Lee, Pei, and Weber (2015); Pei and Yi (2017); Ganong and Jäger (2018)

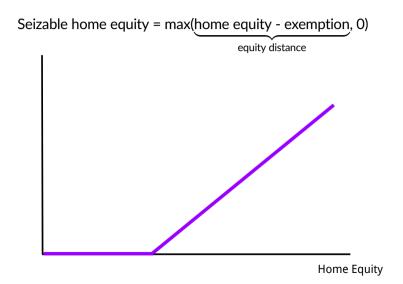
Institutional Background & Data

Determination of Debt Relief Generosity in Bankruptcy

- State asset exemption laws:
 - Limit amount of assets filers can keep in bankruptcy
 - Homestead exemptions protect home equity
- In state s, household i's financial benefit from bankruptcy is:

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Benefit<sub>i,s</sub> = Dischargeable Debt<sub>i</sub> - Seizable Assets<sub>i,s</sub> - Filing Costs<sub>i,s</sub>
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The Homestead Exemption



Main Data: CoreLogic's LLMA

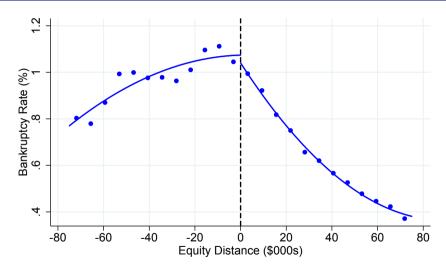
- CoreLogic's Loan-Level Market Analytics (LLMA):
 - Panel of 45% of mortgages originated in the US over 2000-2016
 - Tracks mortgages and bankruptcy filings over time
- Measuring home equity
 - Project initial home value forward over time using Zillow's ZIP-level price index

Empirical Analysis I: The Strategic Bankruptcy Motive

Empirical Strategy: Regression Kink Design (RKD)

- Goal: estimate impact of debt relief generosity in bankruptcy on filing
- Identification Challenges:
 - Unobserved factors affecting both wealth and filing
 - Exclusion restriction (unsecured credit)
- Approach: RKD exploiting kink in seizable home equity
 - ▶ Intuition: est. change in relationship between equity distance and filing at exemption limit
 - ▶ **Key Assumption:** unobs. factors are not kinked functions of equity distance

Kink in Bankruptcy Cost ⇒ Kink in Filing Rate



▶ More Variability

▶ Wider Range

Measurement Error in RKDs

- Imputing home equity ⇒ measurement error
- Measurement error creates non-standard problems for an RKD (sharp and fuzzy)
- New approach: assume curves are quadratic (instead of approx.) within bandwidth
 - ▶ Yields new characterization for bias due to classical measurement error
 - Implies attenuation bias larger when more obs. assigned to wrong side of cutoff
- Implement bias correction using subsample of 200k home sales

Sensitivity of Filing to Financial Cost (RK Results)

Percent change in filings given \$1,000 increase in seizable equity:

	(1)	(2)
	Benchmark	ME-Corrected
RK estimate $\left(\frac{\widehat{\partial p}}{\partial s}\right)$	-1.64***	-3.42***
Title (ds)	1.0	0.42
, ,	(0.21)	(0.44)
Bandwidth	67.07	67.07
Observations	46,026,140	46,026,140

Notes: Coefficients are scaled to correspond to the annual % change in filings per \$1,000 increase in seizable equity. RKD: Optimal bandwidth selection, approximation bias correction, and construction of the robust standard errors follows Calonico, Cattaneo, and Titiunik (2014). Estimation uses a uniform kernel. Statistical significance: 0.05*, 0.01**, 0.001***.

Implies 0.025 percentage point fall in annual filing rate (sample avg: 0.72%)

▶ Various Bandwidths

→ Ganong-Jäger Permutation Test → Heterogeneity

Empirical Analysis II: The Cash-Flow Bankruptcy Motive

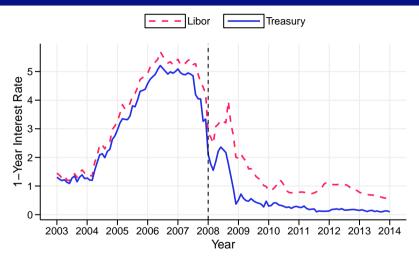
The Cash-Flow Bankruptcy Motive

- Goal: estimate impact of non-seizable cash-flows on bankruptcy filing
- Challenge: finding exogenous shocks to non-seizable resources
 - Non-seizability important for isolating cash-flow motive
- Approach: instrument for mortgage payment reductions
 - Variation comes from adjustable-rate mortgage (ARM) reset rules (similar to Gupta, 2019)
- Mortgage payment reductions
 - Payment reductions not generally seizable in bankruptcy
 - Mortgage debt is not discharged in bankruptcy

Identification: ARM Index

- ARMs (adjustable-rate mortgages)
 - Rate initially fixed (usually +5 years)
 - ▶ New Rate = Pre-Specified Margin + Index Rate
- Popular indexes: 1-year Libor and Treasury rates
 - Libor: daily average interbank loan rate
 - Rate often chosen to match denomination of MBS investors' cost of funds
- Libor and Treasury households have similar mortgage and regional characteristics
 Summary Statistics
 Testing for Differences

Identification: Libor-Treasury Divergence



- ⇒ Treasury-indexed ARMs reset to much lower rates
- Payment difference for median loan peaked at \$4,191 per year

Econometric Specification (IV)

• **Goal:** estimate β , instrumenting for MPay, with IndexRate,

$$B_{ict} = \beta \mathsf{MPay}_i + \alpha_c + \tau_t + \gamma X_{ict} + \epsilon_{ict}$$

- ▶ $B_{ict} = 1$ if household *i* in location *c* files bankruptcy in month *t*
- MPay_i = annual payment component determined by the index rate
- ▶ IndexRate_i = value of i's index rate upon reset
- Expect negative OLS bias: low-risk households can get bigger mortgages
- Exclusion Restriction: index rate only affects filing through payment
- Sample: restrict to 12 months following reset for non-delinquent ARMs

IV Estimation Results

Percent change in filings given \$1,000 increase in mortgage payment:

	(1)	(2)	(3)	(4)
MPay;	30.72***	27.49***	33.49***	29.98***
	(7.36)	(7.64)	(8.48)	(8.71)
	(7.50)	(7.04)	(0.40)	(0.71)
Stage 1 F-Stat.	20.69	18.50	17.11	15.63
Observations	1,092,072	1,092,072	1,092,072	1,092,072
Loan Age FE		√	✓	✓
Loan Age x Time FE			✓	✓
County x Time FE				\checkmark

Notes: All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. Statistical significance: 0.05*, 0.01***, 0.001***.

► 1st Stage ► IV vs. OLS ► Placebo Test

Comparing the Strategic and Cash Flow Motives

Estimate Comparison

- Steps to make cash-flow motive estimate comparable
 - Scale IV estimate to reflect response to change in NPV of mortgage payments
 - Re-weight ARM sample to match RKD on covariates (DiNardo, Fortin, and Lemieux, 1996)
 - ► Yields cash-flow motive estimate: 12.61*** (SE: 3.57)
- \$1,000 reduction in generosity leads to a 3.42% fall in filings
- \$1,000 reduction in payments leads to a 12.61% fall in filings

Theoretical Implications

Model Setting

- Household faces stochastic wealth shocks and incomplete markets
- Has option to file for bankruptcy, decision follows a threshold rule
- Strategic and cash-flow motive estimates corresponds to comparative statics

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Model Setting

- Household faces stochastic wealth shocks and incomplete markets
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- Strategic and cash-flow motive estimates corresponds to comparative statics

Main Result

- $\, \stackrel{cash-flow\ motive}{\rm strategic\ motive} \propto$ decrease in marginal utility when filing (for marginal filer)
- ▶ Relatively stronger cash-flow motive ⇒
 - Marginal filer expects larger consumption increase when filing
 - "Other" costs of bankruptcy are large (e.g., stigma or dynamic costs)
- Note: full info benchmark assumes household observes variation in bankruptcy generosity

Conclusion

Conclusion

- Estimate strength of strategic and cash-flow bankruptcy motives
 - Use RKD and natural experiment
 - ► Find cash-flow motive is 4x stronger than the strategic bankruptcy motive

- Behavior consistent with "other" costs of bankruptcy being large
- Suggests welfare-improving scope for generous bankruptcy

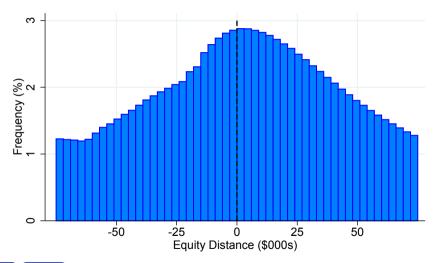
Thanks!

Appendix

Smooth Density Tests

- Test 1: continuous equity distance distribution
 - ► Idea: manipulation around exemption limit ⇒ discontinuous density
 - ► Estimated discontinuity: 1.21% (p-value = 0.24) ✓
- Test 2: smooth predetermined covariates
 - Estimate linear probability model using predetermined covariates
 - Household and mortgage info, ZIP HP growth, county-time FE
 - Test for jump or kink in predicted filing rate
 - ► Estimated kink: -0.04% per \$1,000 (p-value = 0.28) ✓
 - ► Estimated discontinuity: -0.15% (p-value = 0.47) ✓

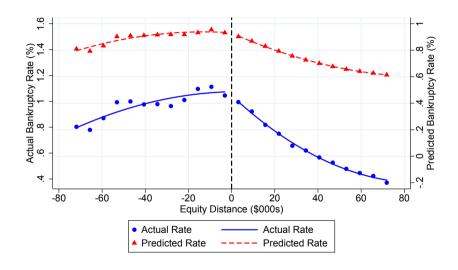
Smooth Density Tests: Equity Distance Distribution



Full Distribution

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Smooth Density Tests: Predicted Filing Rate



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RKD Measurement Error Correction

The measurement-error-corrected estimator is

$$\widehat{\tau}^{PRK-ME} = \frac{\widetilde{\beta}_1^+ - \widetilde{\beta}_1^-}{S'(D)^+ - S'(D)^-}$$

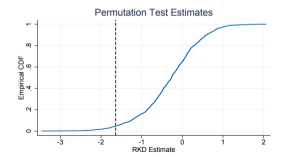
where

$$\widetilde{\beta}_1^+ - \widetilde{\beta}_1^- \equiv \left[\left(1 - \frac{\widehat{\sigma}_{\mu}^2}{\widehat{\sigma}^2} \right) (1 - \widehat{\pi^+} - \widehat{\pi^-}) \right]^{-1} \left(\widehat{\beta}_1^+ - \widehat{\beta}_1^- \right) \xrightarrow{p} (\beta_1^+ - \beta_1^-)$$

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Permutation (Placebo) Test

- Permutation test (Ganong and Jäger, 2018)
 - Conservative, alternative approach to inference
 - ▶ Randomly reassign states' exemption histories and re-estimate RKD
 - ► Compare actual estimate to distribution of 1,000 placebo estimates



Heterogeneity

Percent change in filings given \$1,000 increase:

	Income (ZIP)		Unemp. Ra	Unemp. Rate (County)		Yearly HP Growth (ZIP)	
	Low	High	Low	High	Low	High	
RK Est. $\left(\frac{\widehat{\partial \rho}}{\partial s}\right)$	-2.74***	-2.68***	-1.82***	-1.51***	-2.26***	-0.78*	
()	(0.33)	(0.44)	(0.22)	(0.25)	(0.21)	(0.34)	
Obs. (mil.)	16,586,486	20,063,488	24,898,146	22,740,920	18,672,092	18,737,706	

	Orig.	FICO	Orig	LTV	Predicte	ed P(file)
	Low	High	Low	High	Low	High
RK Est. $\left(\frac{\widehat{\partial p}}{\partial s}\right)$	-2.96***	-0.34+	-1.46***	-2.47***	-0.30*	-1.51***
Obs. (mil.)	(0.35) 19,507,407	(0.18) 18,221,361	(0.23) 23,135,018	(0.26) 24,751,526	(0.12) 12,843,168	(0.31) 11,846,132

Notes: For each covariate I split the sample into two subsets with below and above average values of a single covariate. The coefficients and standard errors are scaled by 1e-8 for readability. Optimal bandwidth selection, bias-correction, and construction of the robust standard errors follows Calonico, Cattaneo, and Titiunik (2014). Estimation uses a uniform kernel. Statistical significance: 0.1⁺, 0.05^{*}, 0.01^{**}, 0.001^{***}.

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RKD: Time Period Splits

	(1)	(2)	(3)
	Pre-Reform	Rush to File	Post-Reform
RK Est.	-1.55***	-7.48***	-1.41***
	(0.37)	(1.53)	(0.20)
Bandwidth	72.64	56.76	70.07
Obs. (mil.)	11.12	1.64	34.51
	Pre-Recession	Recession	Post-Recession
RK Est.	-1.23***	-2.73***	-1.36***
	(0.28)	(0.46)	(0.28)
Bandwidth	72.64	56.760	70.07
Obs. (mil.)	11.12	1.64	34.51

Notes: Each column is the result of estimating the RKD on different sample periods. The pre-reform era is $2000 \, \mathrm{Q1}$ to $2005 \, \mathrm{Q2}$, the rush to file era includes $2005 \, \mathrm{Q3}$ and $\mathrm{Q4}$, and the post-reform era includes $2006 \, \mathrm{Q1}$ to $2016 \, \mathrm{Q1}$. The pre-recession period is defined as $2006 \, \mathrm{Q1}$ to $2007 \, \mathrm{Q4}$, the recession era is $2008 \, \mathrm{Q1}$ to $2010 \, \mathrm{Q4}$, and the post-recession period is $2011 \, \mathrm{Q1}$ to $2016 \, \mathrm{Q1}$. All specification choices match those of the baseline specification. Statistical significance: 0.1+, 0.05^* , 0.01^{**} , and 0.01^{***} .

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IV vs. OLS

Percent change in filings for Libor vs. Treasury-indexed ARMs:

	(1)	(2)	(3)	(4)
		ľ	v	
MPay _i	30.72*** (7.36)	27.49*** (7.64)	33.49*** (8.48)	29.98*** (8.71)
		0	LS	
MPay _i	3.03 (1.87)	2.49 (1.91)	2.46 (1.96)	2.35 (2.23)
Obs.	1,094,998	1,094,998	1,094,998	1,094,998
Loan Age FE Loan Age x Time FE County x Time FE		✓	√ ✓	√

Notes: All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. Statistical significance: 0.05*, 0.01***, 0.001***.

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ARM Placebo Test

Percent change in filings for Libor vs. Treasury-indexed ARMs:

	(1)	(2)	(3)	(4)
Libor _i	3.93 (10.93)	0.25 (10.68)	-0.53 (10.83)	3.70 (11.34)
Obs.	1,094,998	1,094,998	1,094,998	1,094,998
Loan Age FE Loan Age x Time FE County x Time FE		√	√ ✓	√ √ √

Notes: These regressions use data on bankruptcy filings *prior* to the interest rate reset. All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. Statistical significance: 0.05*, 0.01***, 0.001***.

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IV Estimation Results

Percent change in filings given \$1,000 increase in mortgage payment:

	(1)	(2)	(3)	(4)
	2r	nd stage (outcome =	filing increase per \$	1k)
MPay;	30.72***	27.49***	33.49***	29.98***
,,	(7.36)	(7.64)	(8.48)	(8.71)
	1st	stage (outcome = an	nual mortgage paym	nent)
Index Rate _{ict}	1,275***	1,253***	1,384***	1,397***
	(105.97)	(110.08)	(126.52)	(133.54)
Stage 1 F-Stat.	20.69	18.50	17.11	15.63
Observations	1,092,072	1,092,072	1,092,072	1,092,072
Loan Age FE		✓	√	√
Loan Age x Time FE			✓	\checkmark
County x Time FE				✓

Notes: All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. Statistical significance: 0.05*, 0.01***, 0.001***.

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Cash-Flows vs. Expectations: Anticipatory Behavior

Regression using data in 12 months *prior* to reset:

$$B_{ict} = \beta_1 \text{IndexRate}_i + \beta_2 \left(\text{IndexRate}_i \times 2007_t \right) + \alpha_c + \alpha_t + \gamma X_{ict} + \epsilon_{ict}$$

	(1)	(2)	(3)	(4)
IndexRate _i	-0.04 (11.67)	-3.90 (11.56)	-5.90 (11.68)	-1.53 (12.44)
IndexRate $_i$ x 2007 $_t$	-20.38 (12.01)	-19.69 (12.24)	-14.81 (12.05)	-11.56 (14.37)
Observations	1,094,998	1,094,998	1,094,998	1,094,998
Loan Age FE Loan Age x Time FE County x Time FE		√	√ ✓	√ √ √

Notes: I scale coefficient and standard errors on the IndexRate_{ict} covariates so that the coefficient corresponds to the relative (percent) change in the filing rate per 1% increase in the index rate. All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. Statistical significance: 0.05*, 0.01***, 0.001***.

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Estimating Expected Mortgage Payments

• Expected NPV of payments conditional on information in month τ :

$$M_{ au}^{NPV} = \underbrace{s_{ au}(1-\delta)M_{ au}}_{ ext{current payment}} + \underbrace{\sum_{j=1}^{360- au} s_{ au+j}(1-\delta) rac{\mathbb{E}_{ au}(M_{ au+j})]}{1+r}}_{ ext{future payments}}$$

• If households believe payments are martingale ($\mathbb{E}_{\tau}(M_{\tau+j}) = M_{\tau}$) then

$$M_{ au}^{NPV} = M_{ au} \sum_{j=0}^{360- au} rac{s_{ au+j}(1-\delta)}{(1+r)^j} \equiv M_{ au} heta$$

Param.	Value	Meaning	Source
$\{s_t\}_{60\leqslant t\leqslant 360}$	Median: 7 yrs	Survival rate	Estimate using CoreLogic
δ	1.63%	Delinquency rate	Estimate using CoreLogic
r	4.39%	Discount rate	Avg. annual 30-yr FRM rate
θ/12	6.22	Scaling factor	

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Composition Adjustment via Re-weighting

Step 1: Probit Regression

Using RKD and ARM samples, estimate $\hat{p}_i \equiv P(i \in \text{RKD sample})$ as a function of characteristics of interest

Step 2: Construct Weights

$$w_i = \frac{\widehat{p}_i}{1 - \widehat{p}_i} \times \left[\frac{\sum_{i=1}^{N} \mathbf{1}(i \in \mathsf{RKD} \, \mathsf{sample}) / N}{1 - \sum_{i=1}^{N} \mathbf{1}(i \in \mathsf{high} \, \mathsf{group}) / N} \right]$$

Strategic vs. Cash-Flow Motive: Results

		Composition-Adjusted		
		×	✓	
		29.98***	78.45***	
	×	(8.71)	(22.22)	
NPV-Adjusted		4.82***	12.61***	
	/	(1.43)	(3.57)	
Stage 1 F-Stat.		15.63	26.91	

• Representative household lives for two periods

◆ Go Back

• Representative household lives for two periods

◆ Go Back

Consumption when filing for Bankruptcy and Not filing:

$$c_t^B$$

$$c_t^B$$
 , $t=1,2$

$$c_1^N$$
 c_2^N

Representative household lives for two periods

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• Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a$$
 , $t = 1, 2$

$$c_1^N = a$$
 $c_2^N = a$

• a: non-seizable endowment

Representative household lives for two periods

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• Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a$$
 , $t = 1, 2$

$$c_1^N = a + y_1$$
$$c_2^N = a + y_2$$

- a: non-seizable endowment
- *y_t*: stochastic income

Representative household lives for two periods

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Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a + e, \quad t = 1, 2$$

$$c_1^N = a + y_1$$
$$c_2^N = a + y_2$$

- a: non-seizable endowment
- e: exempt assets
- *y_t*: stochastic income

Representative household lives for two periods

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Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a + e, \quad t = 1, 2$$

$$c_1^N = a + y_1 - R_1 d_1 + d_2$$

 $c_2^N = a + y_2 - R_2 d_2$

- a: non-seizable endowment
- e: exempt assets
- *v_t*: stochastic income

- d_t : initial debt
- R_t : gross interest rate

Representative household lives for two periods

◆ Go Back

Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a + e, \quad t = 1, 2$$

$$c_1^N = a + y_1 - R_1 d_1 + d_2$$

 $c_2^N = a + y_2 - R_2 d_2$

Period 1 value functions:

$$V_1^B = u(c_1^B)$$

 $V_1^N(y_1, d_1) = \max_{d_2} u(c_1^N)$

- a: non-seizable endowment
- e: exempt assets
- *y_t*: stochastic income

- d_t : initial debt
- R_t: gross interest rate

Representative household lives for two periods

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Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a + e, \quad t = 1, 2$$

$$c_1^N = a + y_1 - R_1 d_1 + d_2$$

 $c_2^N = a + y_2 - R_2 d_2$

Period 1 value functions:

$$V_1^B=u(c_1^B)-{\color{red}\sigma}$$

$$V_1^N(y_1, d_1) = \max_{d_2} u(c_1^N)$$

- a: non-seizable endowment
- e: exempt assets
- *v_t*: stochastic income
- σ: utility penalty of bankruptcy

- d_t: initial debt
- R_t: gross interest rate

Representative household lives for two periods

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Consumption when filing for Bankruptcy and Not filing:

$$c_t^B = a + e, \quad t = 1, 2$$
 $c_1^N = a + y_1 - R_1 d_1 + d_2$ $c_2^N = a + y_2 - R_2 d_2$

Period 1 value functions:

$$\begin{aligned} V_1^B &= u(c_1^B) - \sigma + \mathbb{E}_1^B \left[V_2^N(y_2, 0) \right] \\ V_1^N(y_1, d_1) &= \max_{d_2} \ u(c_1^N) + \frac{p_2}{p_2} \mathbb{E}^N \left(V_2^B \right) + \frac{1 - p_2}{p_2} \mathbb{E}^N \left[V_2^N(y_2, d_2) \right] \end{aligned}$$

- a: non-seizable endowment
- e: exempt assets
- *v_t*: stochastic income
- σ: utility penalty of bankruptcy

- d_t: initial debt
- R_t : gross interest rate
- p_t: bankruptcy probability

Cash-Flow vs. Strategic Motive

Result: Mapping to Preferences

The relative strength of the cash-flow vs. strategic motive is

$$\frac{-\partial \rho_1/\partial a_1}{\partial \rho_1/\partial e_1} = \frac{u'(c_1^{N\star}) - u'(c_1^B)}{u'(c_1^B)}.$$



Cash-Flow vs. Strategic Motive

Result: Mapping to Preferences

The relative strength of the cash-flow vs. strategic motive is

$$\frac{-\partial \rho_1/\partial a_1}{\partial \rho_1/\partial e_1} = \frac{u'(c_1^{N\star}) - u'(c_1^B)}{u'(c_1^B)}.$$

- Implications a stronger cash-flow motive: $\left(-\frac{\partial p_1}{\partial a_1}>>\frac{\partial p_1}{\partial e_1}\right)$:
 - ▶ Marginal filer's consumption gain is large: $c_1^B >> c_1^{N\star}$
 - But other costs of filing (dynamic costs or stigma) must be large

$$u(c_t^B) - \sigma + \mathbb{E}^B(V_{t+1}) = \max_{d_{t+1}} u(c^{N\star}) + \mathbb{E}^N(V_{t+1})$$

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"Other" Costs of Bankruptcy

Stigma

- ▶ 82% of HHs say default is morally wrong when able to pay (Guiso, Sapienza, and Zingales, 2013)
- ▶ Moral messages can reduce credit card delinquency (Bursztyn, Fiorin, Gottlieb, and Kanz, 2017)

Dynamic credit market costs

- ► Filers see significant increases in credit after flag removal (Musto, 2004; Dobbie et al., 2017; Gross et al., 2018)
- ▶ But insolvent non-filers fare worse in terms of credit access (Albanesi and Nosal, 2018)

Dynamic labor market costs

- ▶ Bankruptcy flags can reduce employment, may signal bad type (Bos, Breza, and Liberman, 2018)
- ► Chapter 13 protection can bolster earnings (Dobbie and Song, 2015)



Bankruptcy Comparative Statics

- Filing decision:
 - Prefer to file when $V_t^B > V_t^N(d_t, y_t)$
 - ▶ Decision follows threshold rule: file if $y_t < y_t^*$
 - ▶ Probability of filing: $p_t = P[y_t < y_t^*(d_t)] = F[y_t^*(d_t)]$
- Effect of a change to period 1's e or a:

$$\underbrace{\frac{\partial p_1}{\partial e_1} = f(y_1^\star) \frac{\partial y_1^\star}{\partial e_1}}_{\text{Strategic Motive}}, \qquad \underbrace{\frac{\partial p_1}{\partial a_1} = f(y_1^\star) \frac{\partial y_1^\star}{\partial a_1}}_{\text{Cash-Flow Motive}}$$

Bankruptcy Comparative Statics: Shifts in the Threshold

• Threshold y_1^* characterized by indifference condition:

$$V_1^B = V_1^N(y_1^*, d_1)$$

Implicitly differentiating the indifference condition yields:

$$\frac{\partial y_1^{\star}}{\partial e_1} = \frac{u'(c_1^B)}{u'(c_1^{N\star})} > 0, \qquad \qquad \frac{\partial y_1^{\star}}{\partial a_1} = \frac{u'(c_1^B) - u'(c_1^{N\star})}{u'(c_1^{N\star})}$$

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Extensions

- Dynamic model: no change
- Allowing filers to save: no change
- Heterogeneity: recast in terms of average marginal filer
- Asset adjustment costs or borrowing constraints: no change if marginal changes in e
 and a don't switch constraints on/off
- Institutional Features:
 - Credit market exclusion for filers: unchanged
 - Delinquency (informal default): unchanged (non-filing MU may correspond to MU in delinquency)

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ARM Summary Statistics

Data: ARMs originated in 2003-2008 (CoreLogic)

		Libor				Treasury	,
	Mean	SD	N	ı	Mean	SD	N
Orig. Bal.	289.2	200.56	51,164	2	54.23	192.4	45,186
Orig. LTV	74.45	12.55	51,164	7	71.24	15.58	45,186
FICO	727.4	47.87	48,237	7	727.2	50.4	43,044
Own. Occ.	82.16		50,423	8	85.93		45,184
UR (county)	9.31	2.51	51,134		9.30	2.49	45,157
Med. Inc (county)	58.91	15.21	51,164	į	59.94	14.69	45,186

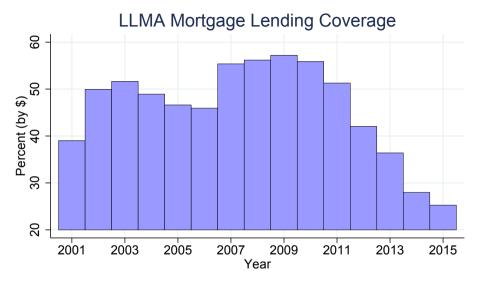
Notes: All values in thousands of 2010 dollars or %.



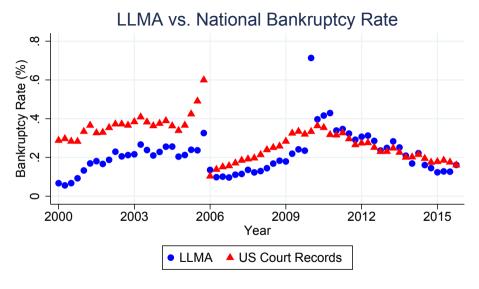
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LHS: Libor Dummy

	(1)	(2)	(3)	(4)
Margin _i	-35.40***	-36.01***	-38.79***	-44.88***
	(1.71)	(1.74)	(2.02)	(3.30)
Old Pay.;	0.04	0.05	0.08	0.03
	(0.05)	(0.05)	(0.05)	(0.06)
Orig. FICO _i	0.001	-0.01**	-0.01**	-0.01
	(4e-3)	(4e-3)	(4e-3)	(0.01)
Orig. LTV;	0.27***	0.22***	0.20***	0.21***
	(0.02)	(0.02)	(0.02)	(0.02)
In(Orig. Bal.);	0.03*	0.01	2e-3	0.02
	(0.01)	(0.01)	(0.01)	(0.02)
UR _{ct} %	-0.29 (0.41)	-0.33 (0.36)	-0.46 (0.38)	
In(Med. Inc.) _{ct}	0.18 (0.22)	0.23 (0.19)	0.24 (0.18)	
$\Delta \ln(HP)_{zt}$	-1e-4	3e-3	3e-3	3e-3
	(4e-3)	(4e-3)	(4e-3)	(0.01)
Observations	61,482	61,482	61,482	61,482
Loan Age FE Loan Age x Time FE County x Time FE		✓	√ ✓	√ √

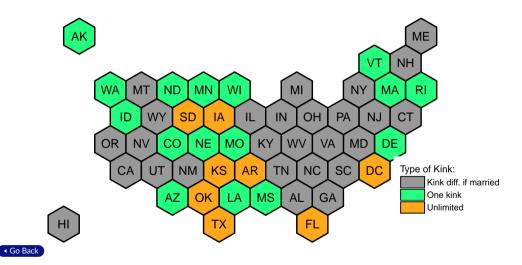






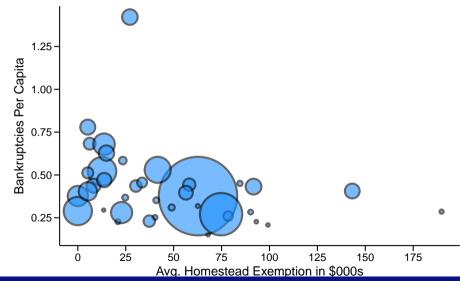


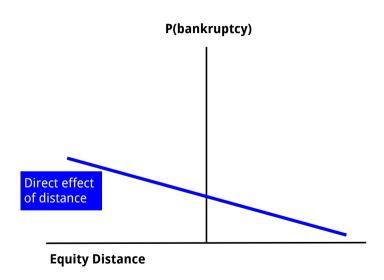
States Used

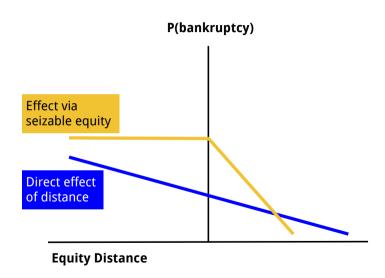


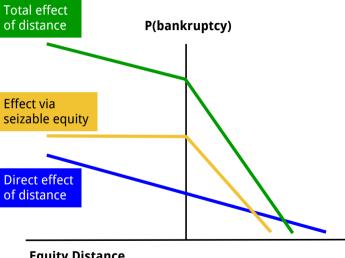
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State Bankruptcy Rate vs. Homestead Exemption

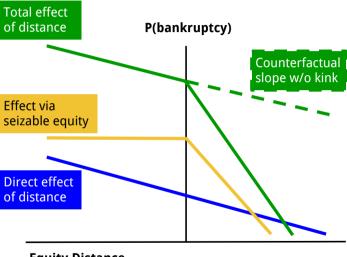








Equity Distance



Equity Distance

Estimand

The RK estimand is

$$\tau := \frac{\partial \mathbb{E}(B|D=0)}{\partial S} = \beta^+ - \beta^-$$

where β^+ and β^- are the RHS and LHS slopes of $\mathbb{E}(B|D)$ at the kink, specifically

$$\beta^+ = \lim_{D_0 \to 0^+} \beta(D_0), \quad \beta^- = \lim_{D_0 \to 0^-} \beta(D_0), \quad \beta(D_0) = \frac{d\mathbb{E}(B|D = \widetilde{D})}{d\widetilde{D}} \Big|_{\widetilde{D} = D_0}$$

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Identification

• Intuition: Let $P \equiv \mathbb{E}[B(D, S(D), u)|D = 0]$. Distance D affects probability P through direct and indirect channels:

$$\frac{dP}{dD} = \frac{\partial P}{\partial D} + \frac{\partial P}{\partial S}S'(D) + \frac{\partial P}{\partial u}u'(D)$$

Differencing the RHS and LHS limits gives

$$\frac{dP}{dD}^{+} - \frac{dP}{dD}^{-} = \frac{\partial P}{\partial S} \left[S'(D)^{+} - S'(D)^{-} \right]$$
=1 (in my case)

if only $\lim_{D\to 0^+} S'(D) \neq \lim_{D\to 0^-} S'(D)$ at D=0



Estimation

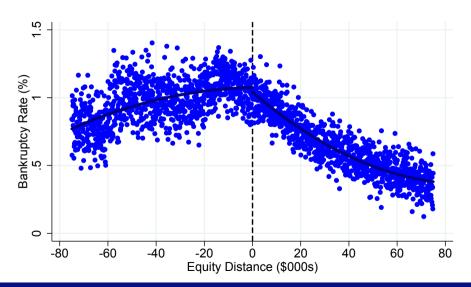
Nonparametric local polynomial estimation:

$$\begin{split} \{\widehat{\boldsymbol{\beta}}_{j}^{+}\} &= \arg\min_{\{\widetilde{\boldsymbol{\beta}}_{j}^{+}\}} \sum_{t} \sum_{i}^{n_{t}^{+}} \left[\boldsymbol{B}_{i,t}^{+} - \sum_{j=0}^{p} \widetilde{\boldsymbol{\beta}}_{j}^{+} (\boldsymbol{D}_{i,t}^{+})^{j} \right]^{2} \boldsymbol{K} \left(\frac{\boldsymbol{D}_{i}^{+}}{h} \right) \\ \{\widehat{\boldsymbol{\beta}}_{j}^{-}\} &= \arg\min_{\{\widetilde{\boldsymbol{\beta}}_{j}^{-}\}} \sum_{t} \sum_{i}^{n_{t}^{-}} \left[\boldsymbol{B}_{i,t}^{-} - \sum_{j=0}^{p} \widetilde{\boldsymbol{\beta}}_{j}^{-} (\boldsymbol{D}_{i,t}^{-})^{j} \right]^{2} \boldsymbol{K} \left(\frac{\boldsymbol{D}_{i}^{-}}{h} \right) \end{split}$$

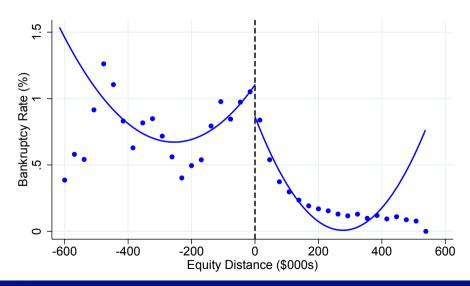
- $B_{i,t} = 1$ if i files for bankruptcy in t
- \triangleright $D_{i,t}$ is i's distance at t from her exemption limit
- $K(\cdot)$ is the kernel and h the bandwidth
- Estimate of interest: $\widehat{\tau} = \widehat{\beta}^+ \widehat{\beta}^-$
- Choose bandwidth to min. MSE (Calonico, Cattaneo, and Titiunik; 2014)

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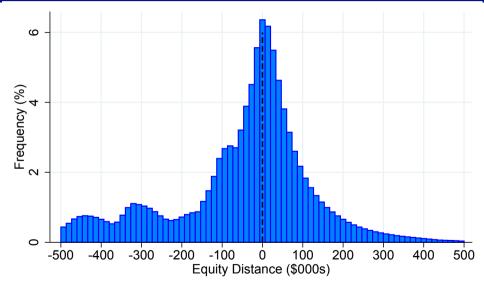
Filing Kink - More Variability



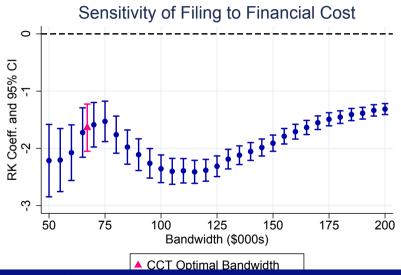
Filing Kink - Wider Range



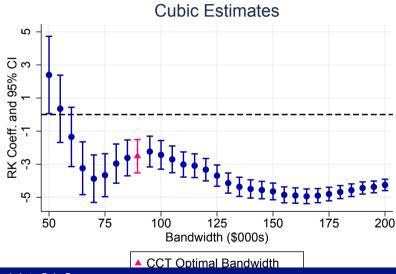
Distribution of Equity Distance to Cutoff



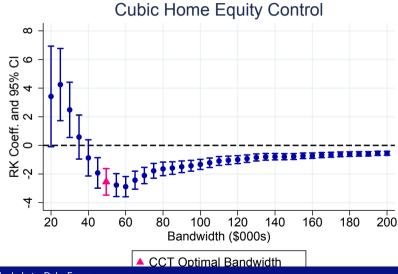
Various Bandwidths (Benchmark)



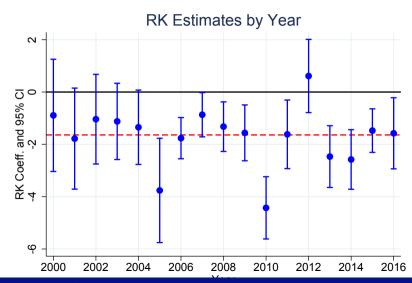
Various Bandwidths (Cubic Equity Distance)



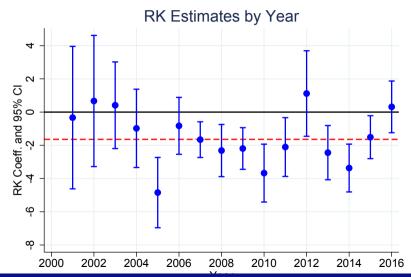
Various Bandwidths (Cubic Home Equity)



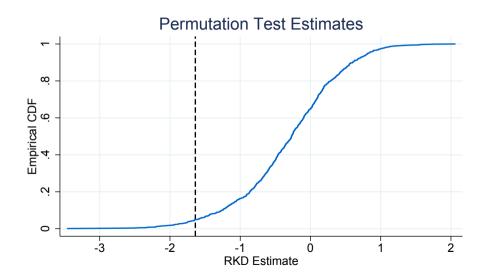
Year-By-Year Estimates



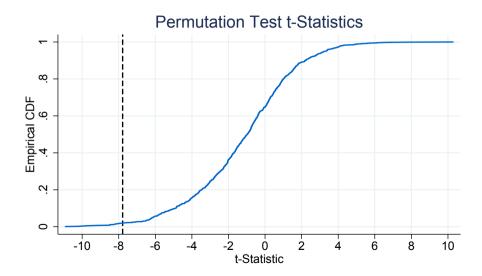
Year-By-Year Estimates (Constant Composition)



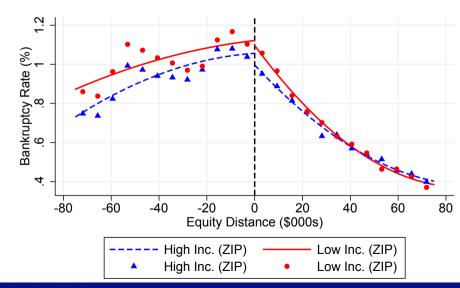
Permutation Test: Coefficient



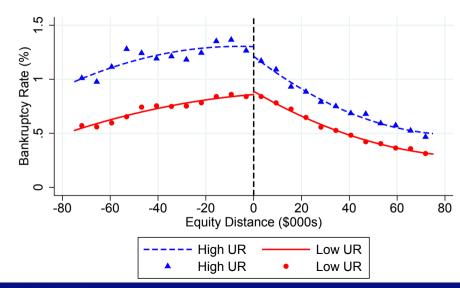
Permutation Test: t-Statistic



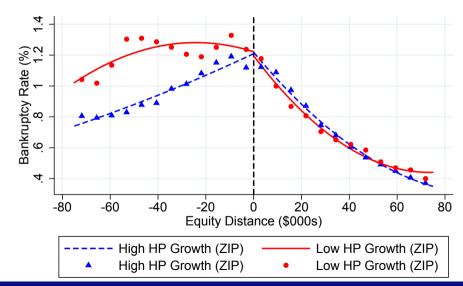
RKD: High vs. Low Income (ZIP)



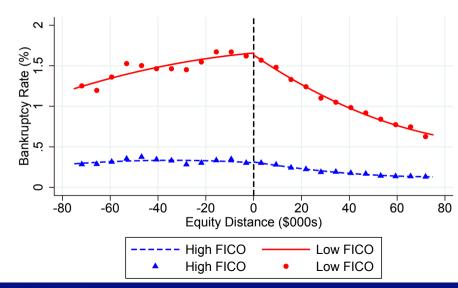
RKD: High vs. Low UR (County)



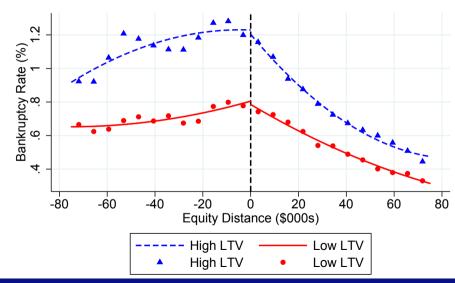
RKD: High vs. Low HP Growth (ZIP)



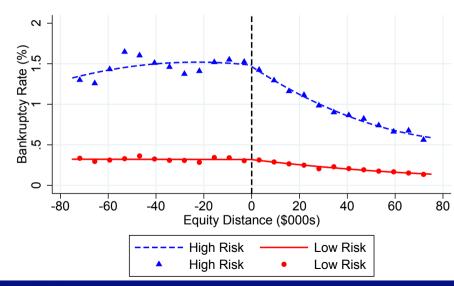
RKD: High vs. Low FICO



RKD: High vs. Low LTV



RKD: High vs. Low Predicted P(file)



RKD: Time Period Splits

	(1) Pre-Reform	(2) Rush to File	(3) Post-Reform	(4) Pre-Rec.	(5) Recession	(6) Post-Rec.		
	Panel A: Unweighted							
Est.	-2.79***	-13.46***	-2.54***	-2.18***	-4.85***	-2.42***		
Std. Err.	(0.67)	(2.75)	(0.36)	(0.50)	(0.81)	(0.49)		
Bandwidth	72.64	56.76	70.07	86.86	66.30	77.13		
Obs. (mil.)	11.12	1.64	34.51	8.63	9.12	18.71		
	Panel B: Weighted for Constant Composition							
Est.	-5.52***	-22.57***	-3.83***	-3.29***	-4.23***	-2.92***		
Std. Err.	(0.98)	(3.59)	(0.45)	(0.70)	(0.84)	(0.66)		
Bandwidth	76.61	55.13	66.04	78.99	78.22	72.89		
Obs. (mil.)	7.83	1.30	29.63	6.78	9.13	16.44		

Notes: These regressions use data on bankruptcy filings *prior* to the interest rate reset. All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. I scale $MPay_{ict}$ and B_{ict} so that the 2^{nd} stage is the effect on the quarterly bankruptcy rate (in percentage points) of a \$10k increase in annual mortgage payments. The 1^{st} stage captures the effect of Libor-indexing on annual mortgage payments. Statistical significance: 0.05^* , 0.01^{***} .

Liquidity Analysis: Restricting to Homeowners

	(1)	(2)	(3)	(4)	(5)	
	2nd stage (outcome = bankruptcy prior to reset)					
MPay _i	0.36**	0.33**	0.35**	0.29*	0.20+	
	(0.12)	(0.11)	(0.11)	(0.11)	(0.10)	
	1st stage (outcome = annual mortgage payment)					
Libor _i	2,086***	2,272***	2,329***	2,390***	2,425***	
	(142.80)	(153.21)	(154.59)	(166.87)	(172.84)	
Stage 1 F	42.68	43.97	45.4	41.03	39.38	
Observations	918,041	918,041	918,041	918,041	918,041	
Loan Age FE		✓	✓	✓	✓	
Loan Age x Time FE			\checkmark	\checkmark	\checkmark	
County x Time FE				\checkmark	\checkmark	
ZIP FE					\checkmark	

Notes: These regressions use data on bankruptcy filings prior to the interest rate reset. All regressions contain county and time FE and household-level controls. Standard errors are clustered by county. I scale $MPay_{ict}$ and B_{ict} so that the 2^{nd} stage is the effect on the quarterly

Baseline Model

- Period 1: ex ante identical households choose borrowing D
- In period 2, households:
 - Realize income $y \sim F_Y$
 - ▶ Receive annuity a, have uncertain illiquid wealth H ~ F_H
 - Can file for bankruptcy and keep e; filing cost φ
- Period 2 consumption with and without filing:

$$c^{NF} = a + y + H - D$$
$$c^{F} = a + e - \varphi$$

Two Bankruptcy Thresholds

• The household prefers to file if

$$c^{NF} < c^{F}$$

$$a + y + H - D < a + e - \varphi$$

$$y < e + D - \varphi - H \equiv y_{MH}^{\star}$$

Household must file if

$$y + a < D$$
$$y < D - a \equiv y_L^*$$

Bankruptcy is driven by liquidity when

$$y_{MH}^{\star} < y_{L}^{\star}$$

 $e - \varphi - H < -a$